
INTERESTINGNESS AS AN INDUCTIVE HEURISTIC FOR FUTURE COMPRESSION PROGRESS

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ABSTRACT

One of the bottlenecks on the way towards recursively self-improving systems is the challenge of *interestingness*: the ability to prospectively identify which tasks or data hold the potential for future progress. We formalize interestingness as an inductive heuristic for future compression progress and investigate its predictability using tools from Kolmogorov Complexity and Algorithmic Statistics. By analyzing complexity-runtime profiles under various priors over computable objects, we demonstrate that the *inductive property of interestingness*—the capacity for past compression progress to signal future discovery—is theoretically viable and empirically supported. Expected future progress depends crucially on the recency of the last observed progress. However, this dependency is highly sensitive to the underlying distribution of objects.

1 INTRODUCTION

A common notion of how a general recursively self-improving open-ended intelligence might be achieved is, in very simple terms, as a cycle between two alternating phases (see Figure 1): In the learning phase, a system uses a learning algorithm—e.g., gradient descent or Reinforcement Learning (RL) techniques—to extract patterns and regularities in the available data, or to gain the skills necessary to solve available problems and tasks. In the generation phase, the insights from the learning phase are used to produce novel artifacts (synthetic data, new tasks, hitherto unsolved problems). From these, the system then can again learn new capabilities, ideally leading to never-ending progress. To achieve true open-endedness, this cycle must be fully autonomous. It cannot rely on human-in-the-loop filtering, hand-crafted curricula, or synthetic data created by researchers.

One example of such a system would be a Large Language Model (LLM) that is iteratively trained on self-generated data, see for example Wang et al. (2023); Zelikman et al. (2024). It is well known that a naive realization of such a setup leads to model collapse (Shumailov et al., 2024; Dohmatob et al., 2024). However, there exist efforts to modify the generation process, or filter the generated data in such a way that continued learning is possible (Lin et al., 2024; Herrmann et al., 2025b;a). Another example would be an RL agent that pursues its own goals in an environment. The goals are chosen according to some intrinsic motivation criterion (Schmidhuber, 1991c; Pathak et al., 2017; Colas et al., 2022).

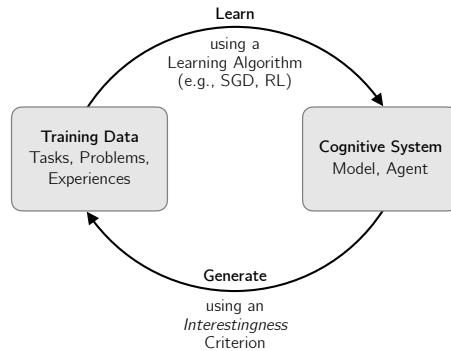


Figure 1: Minimalist depiction of a self-sustaining learning cycle. The Learning phase derives new skills or patterns from data, while the Generation phase creates novel artifacts. To sustain true open-endedness, the generation process must be guided by a criterion that distinguishes learnable structure from noise or already-acquired knowledge.

With the success of large scale models trained with gradient descent and RL techniques, it can be argued that we have a good handle on the learning phase. Developing a generation phase that leads to continued progress, on the other hand, remains an open problem. Using the framing of Hughes et al. (2024), how can we generate artifacts that are both *novel and learnable*? What criterion allows us to select the samples, problems, tasks from which the system can learn something meaningful? How can we distinguish between what is unlearnable, already learned, and so-far unknown but learnable? How can we tell if such an artifact is *interesting*?

To lay out our arguments, this paper proceeds as follows: In Section 2, we review existing interestingness measures, and argue that they are largely post-hoc rather than prospective. This leads us to the question of the predictability of interestingness, and its inductive properties. Section 3 introduces Complexity vs. Runtime profiles. This framework allows us to investigate compression progress over time. In Section 4 we address the induction problem: Can we predict if an object is worth the effort? We analyze how different priors over possible objects—Length, Algorithmic, and Speed—affect our ability to extrapolate future progress. We conclude with a discussion of other kinds of priors that might lead to even more reliable prediction of learnability, and calling for a shift toward *introspective* models that can explicitly assess their own potential for future insight (Section 6).

2 PROPOSED CRITERIA OF INTERESTINGNESS

In the following, we refer to the artifacts a cognitive system (the *subject*) learns from as *objects*. This should highlight the fact that interestingness is not necessarily an intrinsic property of an object in isolation, but a relational property between the object’s structure and the subject’s current internal state.

Count-based, Space Coverage, and Maximum Entropy An object can be deemed interesting if it is *uncommon*, if it has been encountered fewer times than other objects (Sutton, 1990; Bellemare et al., 2016; Tang et al., 2017). In high-dimensional spaces, this requires a density model to quantify novelty via smoothing. If an agent seeks to maximize the entropy of its state-visitation distribution, it will, in the limit, converge toward a policy that encounters all reachable objects with equal probability (Hazan et al., 2019; Mutti, 2023). In certain restricted settings, this is a reasonable measure of interestingness: if it is possible for the cognitive system to encounter all possible objects multiple times, uncommon objects hold the highest surprise in the Shannon sense. But it does not clearly lead to open-ended learning: the system will converge towards the maximum entropy occupancy and then stop learning. If there are more possible objects than can ever be encountered, this measure of interestingness completely depends on some form of coarse-graining or smoothing. In other words, it depends on a *model*.

Prediction Error, Adversarial Criterion This notion moves beyond mere counting by introducing a predictive model with a learning algorithm. Here, interestingness is the inverse of the probability assigned by the model (the “adversarial” criterion). An object that is poorly predicted is considered interesting (Schmidhuber, 1990; 1991a; Pathak et al., 2017). However, this fails to distinguish between epistemic uncertainty (reducible through learning) and aleatoric uncertainty (irreducible randomness). This is the “Noisy TV” problem (Schmidhuber, 2010): a subject using pure prediction error as a reward will get stuck observing a source of pure noise (like a static TV) because it is always unpredictable, yet provides zero learnable structure.

Common approaches to tackle this problem share the insight that an object loses its interestingness once nothing new can be learned about it.

Information Gain, Bayesian Surprise If our model is set up in a probabilistic way, for example as a belief distribution over a set of hypotheses, then we can measure the information gain an object yields. Concretely, this can be quantified as the Kullback-Leibler divergence between the posterior distribution, given the object, and the prior distribution. In more grandiose terms, an object is interesting to the extent it leads to new insights and changes the model’s “world view” (Storck et al., 1995; Itti & Baldi, 2005). Once nothing new can be learned from an object, the information gain subsides. Information gain can be measured not just over a set of fixed concrete hypotheses, but also

for example over latent variables (Tishby & Zaslavsky, 2015; Sekar et al., 2020; Herrmann et al., 2025b) or the predictions of existing data (Herrmann et al., 2025a).

Learning Progress, Competence Progress Closely related to gaining information from an object is the notion of learning progress (Schmidhuber, 1991b; Oudeyer et al., 2007; Stout & Barto, 2010). Instead of framing insight in terms of probabilistic inference and differences between posterior and prior, we can ask: Does the object lead to progress on some pre-defined objective function? An object is then interesting to the extent it improves the model’s performance. Whether this is a reasonable measure of interestingness depends, of course, on the specific choice of objective function. Certainly not all tasks or objective functions are rich enough to allow open-ended learning. However, for many models—especially the ones we might want to use in the setups we mentioned, such as LLMs or world models (Schmidhuber, 2015; Ha & Schmidhuber, 2018; Bruce et al., 2024)—there is one dominant objective function: the negative log-likelihood (NLL). It can be argued that the success of the NLL as a training objective for models is due to the strong connection between learning and compression (Hutter, 2005; Delétang et al., 2023)

Compression Progress This leads us to the concept of interestingness as compression progress (Schmidhuber, 2009; 2006). An object might be deemed interesting if it allows the model to better losslessly compress all available data. For a full compression of the data, we must account for the size of the model itself: what we measure is the number of bits required to describe the model in addition to the bits required to encode the data given the model. This two-part encoding of objects is the core idea behind the Minimum Description Length (MDL) principle (Wallace & Boulton, 1968; Rissanen, 1978). The compression progress associated with an object is thus the reduction of the total encoding size—the sum of the model complexity and the data residual—once the model has incorporated the object. A rich description of compute-bounded and observer-dependent MDL models has recently been presented by Finzi et al. (2026).

The criteria discussed—information gain, learning progress, and compression progress—are closely related in this context: they all measure shifts in the subject’s uncertainty or description length. Information gain provides a probabilistic view of these shifts, learning progress (via NLL) provides an optimization-based view, and compression progress provides an algorithmic view. They provide an *introspective* account, subjective to the system, of how much insight has been gained. While this subjectivity seems at odds with the “objective” measures of Algorithmic Information Theory (AIT) we use later, we view AIT as the theoretical limit that these subjective models strive toward.

All of these criteria implicitly assume that the computational effort has already been spent. In an open-ended setting, this assumption is untenable: the generation phase must decide where to spend effort before the outcome is known. This means an interestingness criterion must be *prospective*. It is not enough to know how much has been learned; we need to know how much we can still learn. The criteria above are largely post-hoc: they quantify the progress made after the computational effort of training has been expended. This makes them insufficient for the “generation phase,” where the system must select promising objects from a vast space of possibilities before committing significant resources to them.

In everyday speech, we often call a thing interesting because we have a feeling that it holds further secrets. This is the sense of “interesting” we advocate for: a prospective heuristic that predicts future learning progress based on past experience. This leads to a fundamental question: under what conditions is such a prediction even possible? To investigate this, we can formalize the relationship between past and future progress through the lens of complexity and runtime. It should be mentioned that many more conceptions of interestingness have been proposed. We discuss two further major directions in the Appendix: notions based on social or multi-agent interactions (A.1), and LLMs emulating a human-like judgement of interestingness (A.2).

2.1 CAN WE LEARN HOW TO PREDICT LEARNABILITY?

Before moving to our formal investigation, we must consider whether the prediction of learnability can be treated as just another learning problem. As described, information gain, learning progress, and compression progress measures require running the inference or learning algorithm to evaluate an object. This makes them not directly usable in an open-ended setting: we cannot simply “train on everything” to see what works. Doing so is not only computationally prohibitive but also risks

model collapse, as we have mentioned before. Instead, we want to *predict* these measures before investing the effort to learn from them. What we usually do when we have “post-hoc” data and want to predict it is *learning* how to predict it. This is indeed what happens when compression progress is used as an intrinsic reward for an RL agent, such as in a Controller-Model setup (Schmidhuber, 2015; Kompella et al., 2017).

However, a meta-cognitive property like interestingness can be difficult to learn for several reasons:

- **Non-stationarity**—Interestingness is a moving target. As the subject learns, an object that was once insightful becomes boring, requiring the meta-model to constantly adapt to the subject’s changing state.
- **Data Sparsity and Cost**—Ground-truth labels for progress are expensive to obtain, as they might require executing a full training phase just to evaluate a single data point or task.
- **Credit Assignment**—In the many batches of data required for training, it is profoundly difficult to identify which specific samples contributed to the reduction in loss. This credit assignment problem makes it hard for a predictor to associate specific object features with the resulting learning progress.

We have made informal arguments why it is difficult to predict future learnability. Essentially, we are asking about the *inductive property of interestingness*: can we infer how much there is yet to learn—and how easily accessible that content is—based on the trajectory of how much we have already learned? To answer this rigorously, we now turn to a formal setting where learning progress is mapped to algorithmic complexity and runtime.

3 RUNTIME AND COMPLEXITY PROFILES

To analyze interestingness with sufficient generality, we move to an abstract setting using the tools of Algorithmic Information Theory. We consider all data and artifacts as binary strings $x \in \{0, 1\}^*$. Let $K(x)$ denote the standard prefix-free Kolmogorov complexity of x : the length of the shortest program p on a universal prefix Turing machine that halts and outputs x . In the context of open-ended learning, we are not merely interested in the absolute compression limit $K(x)$, but in the compression progress made as a function of computational effort. Let $K^r(x)$ be the time-bounded Kolmogorov complexity: the length of the shortest program that computes x within r steps. While $K^r(x)$ is machine-dependent, it allows us to formalize the learning process as a trajectory of decreasing description lengths over time.

Our system has a history of experiences, partially compressed it into some form B . This representation B is the current state of our model. Now we observe a new object o . Introduction of a new object is concatenation. By this concatenation, we’re asking “is o interesting in the current context?”, instead of “is o interesting, in and of itself?”. This is exactly the relational property between model and object we discussed at beginning of Section 2.

Now three qualitatively different things can happen. First, o could be algorithmically random relative to B —it carries no learnable regularities, and the compression of Bo behaves as if B alone were being further compressed, with o contributing only irreducible noise. Second, o could be itself compressible—it contains internal regularities that a short program can exploit, yielding immediate but local compression progress. Third, and most interestingly, o and B could be *relationally* compressible: o contains structure that unlocks new descriptions of B itself, or vice versa— B provides context that makes o dramatically shorter to describe. This third case is precisely what we mean by o being interesting in context. It could of course also be some mixture of these three cases. The point is, the interestingness of object o , in this framing, is measured by the expected compression progress of Bo . For our theoretical analysis, we simply define $x := Do$. Since we want a *prospective* measure—one that predicts future compression before it is achieved—we study not the final compression $K(Do)$, but the trajectory $r \mapsto K^r(Do)$: how description length decreases as more computational effort is allowed. A drop in this profile at runtime r represents a compression breakthrough achievable with effort r . Interestingness is then the question of whether this profile, observed up to a current runtime budget, will continue to drop for larger r .

Complexity vs. Runtime Profiles To study this discovery process, we use the complexity vs. runtime profile (Vereshchagin & Shen, 2016; Bauwens, 2010; Antunes et al., 2017):

$$D_x = \{(r, c) \mid K^r(x) \leq c\}.$$

The boundary of this profile, $c(r) = \min\{c \mid (r, c) \in D_x\}$, tracks the shortest description of x available given a runtime budget r . An object is *logically deep* (Bennett, 1988) if this boundary continues to drop significantly even for very large r . Our central question is about the inductive properties of the profile: Does the shape of D_x for $r \leq R$ allow us to predict the existence of further “drops” in complexity for $r > R$?

Log-Size vs. Complexity Profile and Sophistication To leverage results from algorithmic statistics, we introduce a related profile: the log-size vs. complexity profile P_x :

$$P_x = \{(i, j) \mid \exists A \text{ s.t. } x \in A, K(A) \leq i, \log \#A \leq j\},$$

where $A \subseteq \{0, 1\}^*$ is a computable finite set of strings. The boundary of P_x represents, for a given set complexity, the optimal two-part description consisting of a model (the set A) and the data-to-model code (the index of x within A). This profile always has a negative slope of at most -1 , reflecting the trivial exchange between halving the set size at the cost of adding one bit to the description length of the set. A “drop” in P_x occurs when the boundary falls below the line of slope -1 . Such drops indicate that the corresponding model contains additional structural information about the string, beyond the random information of the index. The point where the boundary last meets the -1 slope line before following it indefinitely defines the *sophistication* of x (denoted m_x), representing the complexity of the “best” model for x (Koppel, 1987; Antunes & Fortnow, 2009).

The $P_x \leftrightarrow D_x$ Correspondence The relationship between complexity vs. runtime and log-size vs. complexity profiles is bridged by the Busy Beaver function $\text{BB}(k)$ —the maximum number of steps a halting k -bit program can run. While BB is uncomputable, it serves as a universal time-scale that abstracts away machine dependence. With this rescaling, P_x and D_x are approximately affine transforms of each other:

$$(i, j) \mapsto (\text{BB}(i), i + j), \quad (1)$$

within logarithmic precision $O(\log |x|)$ (see Vereshchagin & Shen (2016), Theorems 4 and 6). This correspondence is profound: it implies that every drop in P_x (a new non-trivial model with a higher complexity) corresponds to a drop in D_x (a progress in compression when the maximum runtime is increased). Figure 2 illustrates the connection between the two profiles.

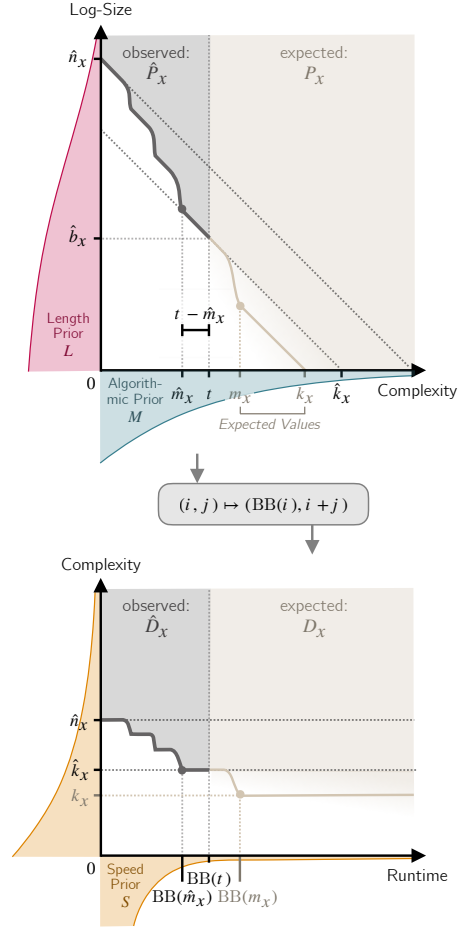


Figure 2: Complexity and Profile Dynamics. (Top) The log-size vs. complexity profile (P_x) of a string x , with the partial profile (\hat{P}_x) observed up to complexity t . The observed characteristic values are \hat{m}_x (complexity of the last drop) and \hat{b}_x (log-size of the smallest set at complexity t), which determines the current estimated complexity \hat{k}_x . We evaluate the expected continuation of \hat{P}_x under various priors; the expected values $\mathbb{E}[m_x]$ and $\mathbb{E}[k_x]$ are functions of the gap $t - \hat{m}_x$, representing the effort expended since the last discovery. (Bottom) The complexity vs. runtime profile (D_x) derived from P_x via the transformation in Eq. 1. Priors are indicated on their relevant axes: the **Length Prior** (string length at complexity 0), the **Algorithmic Prior** (program complexity), and the **Speed Prior** (the product of distributions over complexity and log-runtime, as sketched on the lower axes).

3.1 HOW PROFILES RELATE TO CURRENT MACHINE LEARNING

Neural Networks & SGD Assume we train a neural network with SGD on some data x using an NLL objective. We can plot the trajectory of training in the (i, j) space of the P_x profile, with the effective complexity of the weight function on the x-axis and the log-size of the set of inputs assigned non-negligible probability on the y-axis. At random initialization, the network computes a near-uniform function over inputs—high log-set-size. The weights have initially very low complexity: All we need in order to describe the weights is the random seed and the random number generator. Also the *functional* complexity of this map is low, since it can be described simply as “approximately uniform.” As SGD proceeds, the function becomes more specific: log-set-size decreases as the model assigns higher probability to training data, and model complexity increases as the learned map encodes more structure, and the weights become more refined. If the network merely memorizes the data, this progression follows the trivial slope of -1 : each additional bit of weight complexity eliminates roughly half the plausible candidates, with no net compression gain. A genuine structural discovery—for example the network finding a regularity that explains many examples at once—corresponds to a *drop* below this slope: description length falls faster than one bit per bit of added complexity. This is exactly a drop in the P_x profile, and in MDL terms, it is the moment the two-part code (weights + residual) decreases. In practice, SGD instantiates a single trajectory through this space, not necessarily the optimal boundary of P_x , but the qualitative structure—periods of slope- -1 memorization punctuated by drops corresponding to genuine discoveries—should be preserved.

The same story can be told in the D_x framing, where the connection to effort is more direct. Here the y-axis is total description length—weight complexity plus NLL—and the x-axis is runtime, meaning the number of gradient steps taken. Each drop in this profile represents a point where additional compute yielded a genuine compression gain rather than mere redistribution of bits between model and residual.

Chain-of-Thought and Recurrent Reasoning Chain-of-Thought maps naturally onto the D_x profile. Each reasoning step extends the computation used to produce the final answer, so reasoning steps are counted on the x-axis. The y-axis is the description length of the answer *given* the reasoning trace so far: how many bits are needed to specify the correct answer once the trace is fixed. A drop in this profile occurs where longer, more deliberate, more effortful reasoning leads to higher likelihood of the target. The same framing applies directly to depth-wise recurrent architectures, where each recurrence step plays the role of a reasoning token: runtime depth is the x-axis, and a drop corresponds to a recurrence step that increases the likelihood of the correct output.

Model Scaling Model size offers a third axis along which effort can be scaled. When models of increasing parameter count are each trained for a fixed number of SGD steps, they trace a trajectory in the P_x profile: parameter count is a proxy for the complexity of the model (x-axis), and the NLL on training data is the log-size of the set of plausible inputs (y-axis). A drop in this profile occurs when a larger model finds a description of the data that is *shorter than the slope- -1 extrapolation* from smaller models—that is, when increased capacity enables a genuine structural discovery rather than merely additional memorization.

Alternatively, training FLOPs provide a more direct correspondence to the D_x profile, where runtime is the x-axis and total description length—weight complexity plus NLL—is the y-axis. Here, a drop occurs when additional compute yields genuine compression of the full system, not merely redistribution between model size and residual loss. This total description length can be approximated empirically by the area under the training loss curve, via pre-quential coding: encoding each training example sequentially given the model at that step accumulates the per-step description lengths, approximating the MDL two-part cost.

3.2 COUNTING STRINGS WITH SPECIFIC PROFILES

Any log-size vs. complexity profile P has three characteristic values: The log-size value of the leftmost boundary point $n_P = \min\{r \mid (0, r) \in P\}$ corresponds to the log length of a string with this profile. The complexity value of the rightmost boundary point $k_P = \min\{r \mid (r, 0) \in P\}$ corresponds to the Kolmogorov complexity of a string with this profile. And the complexity

value of the point where the boundary meets the diagonal line leading to $(k_P, 0)$, namely $m_P = \min\{(r, k_P - r) \in P\}$, corresponds to the sophistication of a string following this profile.

We can quantify how many strings have a profile close to P , where P is a valid profile, meaning an upward-closed set whose boundary has a slope of at most -1 , or formally, for which $(a, b + c) \in P \implies (a + b, c) \in P \forall a, b, c$. Vereshchagin & Shen (2016) (Theorem 19) limit the minimum number of strings close to P :

$$\#\{x \mid P_x \approx P\} \geq 2^{k_P - m_P + O(1)}, \quad (2)$$

where the closeness is of order $O(C(P) + \log n_P)$. Theorem 20 of the same work proves the maximum number of strings following this profile: There are at most

$$\#\{x \mid P_x \approx P\} \leq 2^{k_P - m_P(\epsilon) + 2\epsilon + O(\log n_P)} \quad (3)$$

strings that are ϵ close to profile P . The value of $m_P(\epsilon)$ is similar to m_P .

The first important observation is that for any given valid profile P , there exist strings which have a profile close to it. The term $k_P - m_P$ represents the “random” part of the string—the bits that cannot be further compressed into a structural model. Crucially, the number of strings with a given profile is dominated by this residual randomness. This suggests that while structure is rare, strings with the *same* structural profile are plentiful, differing only in their structureless noise. The value n_P , i.e. the log-length associated with profile P , contributes only logarithmically, compared to k_P and m_P , to the upper bound of the number of strings. The exact shape of the profile—beyond the values n_P, k_P and m_P —only plays a role via its complexity $C(P)$ in the closeness to the lower bound of strings. In the next section, we use these bounds to determine whether observing a partial profile \hat{P}_x up to complexity t allows us to extrapolate the full profile.

4 QUANTIFYING THE INDUCTIVITY OF INTERESTINGNESS

If we observe a *partial profile* \hat{P}_x up to (but not including) complexity t , what can we infer about its continuation? Operationally, the boundary of \hat{P}_x corresponds to the best compression achieved so far under a fixed compute budget (via $P_x \leftrightarrow D_x$), not to full access to the profile. Specifically, we look at the expectations for the last drop (m_x) and the ultimate complexity (k_x). To formulate these expectations, we must assume a prior distribution over strings, representing the “world” our learning system inhabits. The concepts from this section are summarized in Figure 2.

4.1 LENGTH PRIOR

The simplest prior considers only string length, effectively asking for the probability that a monkey randomly typing bits generates x . To ensure this is a valid semi-measure, we use a prefix-free encoding (e.g., duplicating bits and ending with 01),

$$L(x) = 2^{-(2|x|+2)}. \quad (4)$$

Using the bounds from Equations (2) and (3), we find that for strings sampled from L , the observed partial profile is highly predictive:

Proposition 4.1. *Let \hat{P} be an observed partial profile up to (but not including) complexity t , with the last observed drop at $(\hat{n}, \hat{k} - \hat{n})$. Assuming the object producing \hat{P} is sampled from the Length Prior L , the following properties hold:*

(i) *The probability, given \hat{P} , of a further drop in the profile at complexity $m \geq t$ is:*

$$p_L(m \geq t \mid \hat{P}) = 2^{-(t - \hat{m}) + O(\log \hat{n})}.$$

(ii) *Given the existence of such a drop and assuming $\hat{k} - t \gg 0$, the expected final complexity is:*

$$\mathbb{E}_{x \sim L}[m \mid \hat{P}, m \geq t] \approx \hat{k} - 2.$$

(iii) Without conditioning on a further drop, the expected further compression progress is:

$$\hat{k} - \mathbb{E}_{x \sim L}[k \mid \hat{P}] = 2^{-(t-\hat{m})+O(\log \hat{n})}.$$

The proof can be found in Appendix B.1. We can see that the difference $t - \hat{m}$ determines the probability of further compression after t , and the amount of expected further compression progress. Before we discuss the implications, we look at two more reasonable priors over strings.

4.2 ALGORITHMIC PRIOR

We can explain the Algorithmic or Solomonoff prior (Solomonoff, 1964) in similar terms as the length prior. But instead of the monkey typing the string directly, we now take the probability of the monkey typing a program that, when run on a prefix Turing machine U , outputs x . The Algorithmic prior is

$$M(x) = \sum_{p:U(p)=x} 2^{-|p|}. \quad (5)$$

This prior prefers algorithmically simple strings (ones with a short description length) as opposed to the Length Prior, which prefers short strings.

Proposition 4.2. *Assuming the object producing \hat{P} is sampled from the Algorithmic Prior M rather than the Length Prior, the following properties hold:*

(i) *The probability of a further drop at or after complexity t is:*

$$p_M(m \geq t \mid \hat{P}) = \frac{1}{1 + \frac{2^{-\hat{m}+O(\log \hat{n})}}{2^{-t+O(\log \hat{n})} + 2^{-\hat{k}+O(\log \hat{n})}}},$$

which, for $\hat{k} - t \gg 0$, simplifies to:

$$p_M(m \geq t \mid \hat{P}) \approx 2^{-(t-\hat{m})+O(\log \hat{n})}.$$

(ii) *Given the existence of such a drop, the expected final complexity is:*

$$\mathbb{E}_{x \sim M}[m \mid \hat{P}, m \geq t] \approx \frac{\hat{k} + t}{2}.$$

(iii) *Without conditioning on the existence of a further drop, the expected further compression progress is:*

$$\hat{k} - \mathbb{E}_{x \sim M}[k \mid \hat{P}] \approx 2^{-(t-\hat{m})+O(\log \hat{n})}.$$

For the proof, please see Appendix B.2. For the Algorithmic Prior, further compressibility is equally determined exponentially by $t - \hat{m}$. However, we can show that nevertheless, the Algorithmic Prior yields higher expected compression.

Proposition 4.3. *Assuming $\hat{k} - t \gg 0$ and $t - \hat{m} \gg \log \hat{n}$, the Algorithmic Prior M yields significantly higher expectations for future insight than the Length Prior L . Specifically, the following ratios hold:*

(i) *The ratio between the further drop probabilities is approximately linear in the remaining complexity gap:*

$$\frac{p_M(m \geq t \mid \hat{P})}{p_L(m \geq t \mid \hat{P})} \approx \hat{k} - t - 1.$$

(ii) *The ratio between the expected further compression progress is approximately quadratic in the remaining complexity gap:*

$$\frac{\hat{k} - \mathbb{E}_{x \sim M}[k \mid \hat{P}]}{\hat{k} - \mathbb{E}_{x \sim L}[k \mid \hat{P}]} \approx \frac{1}{4} (\hat{k} - t - 1) (\hat{k} - t).$$

The proof can be found in Appendix B.3. These ratios demonstrate that the inductive property of interestingness—the capacity for past compression progress to signal future discovery—is significantly more robust under the assumption of algorithmic regularity.

While our primary analysis considers the indefinite future, we show in Appendix B.1.5 and ?? that the probability of discovery is heavily concentrated near the current observation threshold. For any finite computational window, the likelihood of a ‘breakthrough’ vanishes exponentially as the gap since the last progress increases, providing a rigorous justification for prioritizing tasks with the most recent ‘aha!’ moments.

4.3 SPEED PRIOR

The algorithmic prior M is uncomputable, and it is impossible to obtain exact samples from it. We can see this from the fact that the program that the monkey types might not halt—and it is in general not possible to determine whether that is the case. The Speed Prior $S(x)$ (Schmidhuber, 2002b) addresses this by penalizing both program length and runtime. Intuitively, for the Speed Prior, the monkey has to first spell out the maximum runtime in a prefix-free format, and then the program computing x . If the program has not halted before the specified runtime, the output is assumed to be the empty string. We write $p \rightarrow_i x$ iff it takes $2^{i-|p|}$ steps for program p to compute output x . The speed prior is defined as

$$S(x) = \sum_{i=1}^{\infty} \sum_{p \rightarrow_i x} 2^{-(i+|p|)}. \quad (6)$$

It is closely related to Levin Complexity $Kt(x) = \min\{|p| + \log \text{time}(p) : U(p) = x\}$ (Levin, 1984).

Proposition 4.4. *Assuming the object producing \hat{P} is sampled from the Speed Prior S , the following properties hold:*

(i) *The probability of a further drop at or after complexity t is:*

$$p_S(m \geq t \mid \hat{P}) \approx 0.$$

(ii) *Given the existence of such a drop, the expected final complexity is:*

$$\mathbb{E}_{x \sim S}[m \mid \hat{P}, m \geq t] \approx \frac{\hat{k} + t - 1}{2}.$$

(iii) *Without conditioning on the existence of a further drop, the expected further compression progress is:*

$$\hat{k} - \mathbb{E}_{x \sim S}[k \mid \hat{P}] \approx 0.$$

The proof can be found in Appendix B.4. The Speed Prior is the most ‘conservative’. Because it penalizes runtime, it assumes that if a faster way to compress the string existed, it would have been found already. This prior essentially treats the current frontier of compression as the final one.

4.4 DISCUSSION: THE INDUCTIVE PROPERTY OF INTERESTINGNESS

With Equation (1), it is possible to transform the above results from P_x profiles to the compression against effort profile D_x , since the strings—and hence also the numbers of strings—with a certain

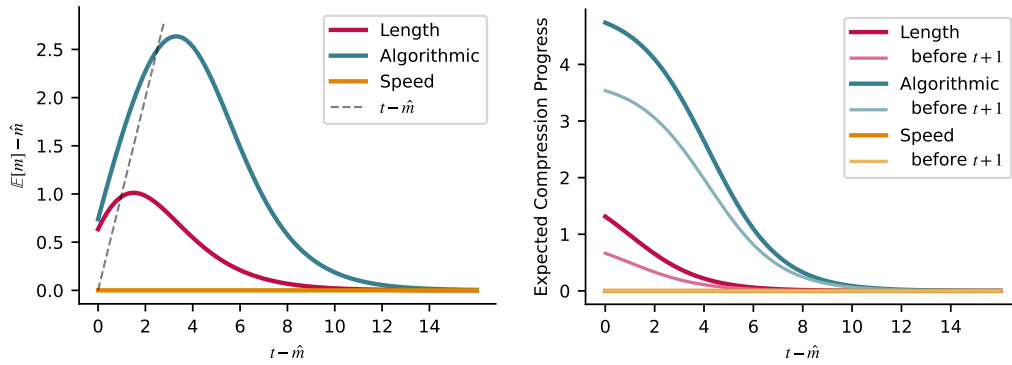


Figure 3: The expected difference between the complexity of the last drop m and the last observed drop \hat{m} (left), and the expected future compression progress (right). Both are plotted as a function of the difference $t - \hat{m}$, for the Length, Algorithmic, and Speed Prior.

P profile shapes are the same as the ones with the corresponding D profile shapes. Thus, we see that under standard priors, the “Inductive Property of Interestingness” holds a specific mathematical form: past compression progress can be an indicator of future progress, but only if that progress is recent relative to the computational effort expended. As shown in Figure 3, if the gap $t - \hat{m}$ is small, there is an expectation for future drops, at least for the Length and Algorithmic Prior. This may justify investing resources into objects that have recently yielded insight. Conversely, if a large computational gap exists without progress, the probability of a future “aha!” moment vanishes exponentially. Note, however, that this does not mean that it is impossible: as mentioned in Section 3.2, all valid profile shapes are closely followed by some strings. They are just rare under the priors we investigated. But in general, all else being equal, the most promising object for further engagement is the one which showed the most recent progress.

A critical caveat of the $P_x \leftrightarrow D_x$ correspondence is its use of the Busy Beaver function (BB). Because BB grows faster than any computable function, the runtimes discussed here are physically unrealizable. However, BB serves a vital theoretical role: it abstracts away machine dependence, ensuring these results reflect the *intrinsic* algorithmic nature of the objects rather than the specifics of a reference computer. This machine independence is important, since there cannot be any particular “most basic” or “most universal” computer (Müller, 2010). While these findings are theoretical, they provide some conceptual grounding. They suggest that “interestingness” can be a principled heuristic for navigating the space of computable objects. To corroborate our findings, we now test the inductive property of interestingness experimentally in three different computationally universal paradigms.

5 EMPIRICAL RESULTS

The theoretical results from the previous section hold only for the busy beaver regime, which involves runtimes that are not physically realizable. While they provide insight into the fundamental algorithmic properties of objects and their time-bounded compressibility, it is not immediately obvious how relevant these results are to real-world scenarios. To address this, we conduct empirical experiments that mirror our theoretical results. We choose three fundamentally different, yet still universal and practical, computational paradigms that represent a wide variety of possible universal computers: 2-Tag systems, elementary cellular automata with Rule 110, and the Brainfuck (BF) language.

In each of these systems, we run all programs up to a certain length with generous runtime limits. This allows us to construct real-world runtime vs. complexity profiles for all objects (i.e., output strings) computed by multiple different programs. From these profiles, we can compute the relationship between steps since the last observed progress and either the empirically achieved further compression progress or the relative time of the last further compression progress. We re-weight

objects based on their prior probabilities and plot the curves in Figures 4, 5, and 6. These are the empirical equivalents of the theoretical expectations shown in Figure 3.

We observe that in all three computational paradigms, the fundamental trends clearly hold: Recent compression progress implies further compression progress in the future, especially under the assumption of the Algorithmic Prior. The most significant difference from the theoretical results is with respect to the Speed Prior: in the heavily constrained empirical setting, the logarithmic bias against long runtimes is less pronounced. This is especially the case for Rule 110 cellular automata, for which—due to their construction—no programs with very short runtimes exist. The exact experimental setups for the three systems are detailed below.

5.1 2-TAG

In this system, we run 2-tag systems (Post, 1943) with the alphabet $\{a, b, c, H\}$, where H is the halting symbol. These kinds of systems have been shown to be Turing complete (Wang, 1963; Cocke & Minsky, 1964). We enumerate all possible production rules up to a combined length of 13, with H appearing only at the end of a single rule. The starting word is always ‘aaa’. This leads to approximately 120 million programs which we run for up to 100k steps. If no output is produced within this limit, we consider the program non-halting. For the results shown in Figure 4, we filter out all objects that show no drops in the runtime vs. complexity profile, as they do not meaningfully contribute to the analysis.

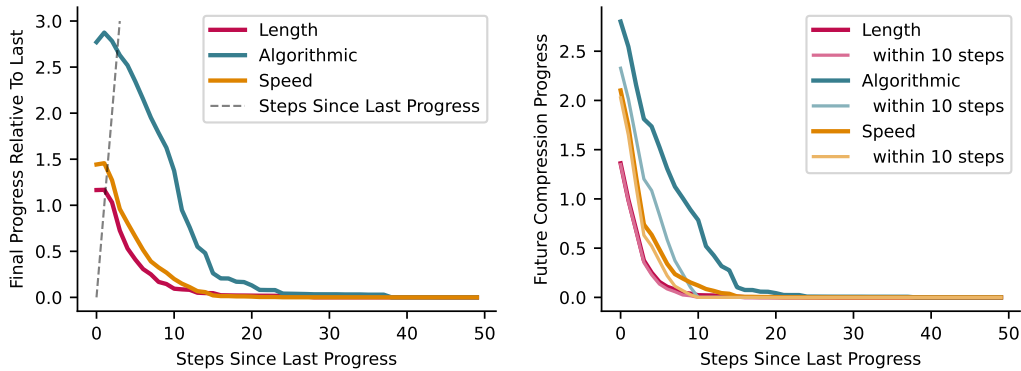


Figure 4: 2-Tag Systems. The average step of the final compression progress (left), and the average remaining compression progress (right), relative to the most recent progress. Plotted as a function of the steps since this most recent progress. As the number of these steps increases, the average remaining compression progress vanishes. Curves shown for outputs weighted with the Length, Algorithmic and Speed Prior. These results mirror the theoretical findings from Section 4.

5.2 RULE 110

Rule 110 is a Turing-complete elementary cellular automaton (Wolfram, 1983; Cook et al., 2004). We use a state of size 512 with a cyclic boundary condition¹. The programs are defined as the leftmost bits of the tape, up to the last value of 1. We enumerate all programs up to length 25, resulting in approximately 34 million simulations. A program halts as soon as any particular state appears for a second time; this state is then the output of the program. If no state repeats within 100k steps, we consider the program non-halting.

5.3 BF

BF (Müller, 1993) is a highly compact Turing-complete programming language. Since we are not providing any external inputs, we do not use the *read* instruction ‘,’. We enumerate all programs up

¹We do this purely for practical reasons and are aware that this technically reduces the automaton to a finite state machine.

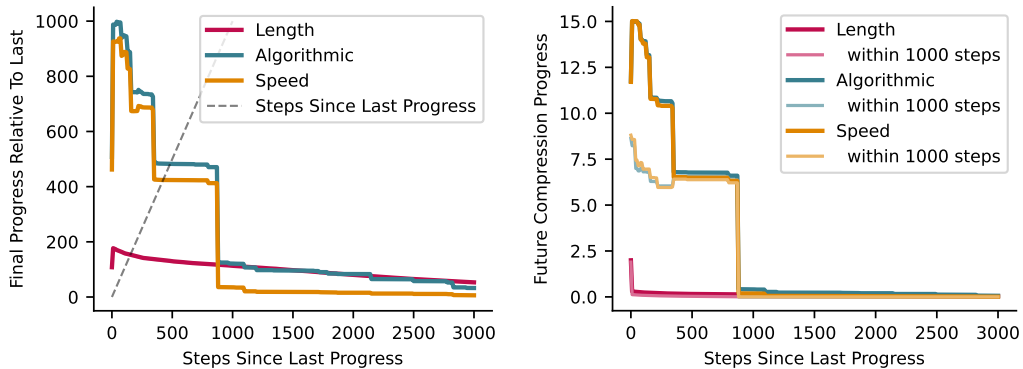


Figure 5: Rule 110. Analogous to Figure 4. Due to there not being any programs with very short runtimes, the difference between Algorithmic Prior and Speed Prior is less pronounced.

to length 11 consisting of the remaining 7 instructions, leading to approximately 2.3 billion different programs, which we run for up to 100k steps.v

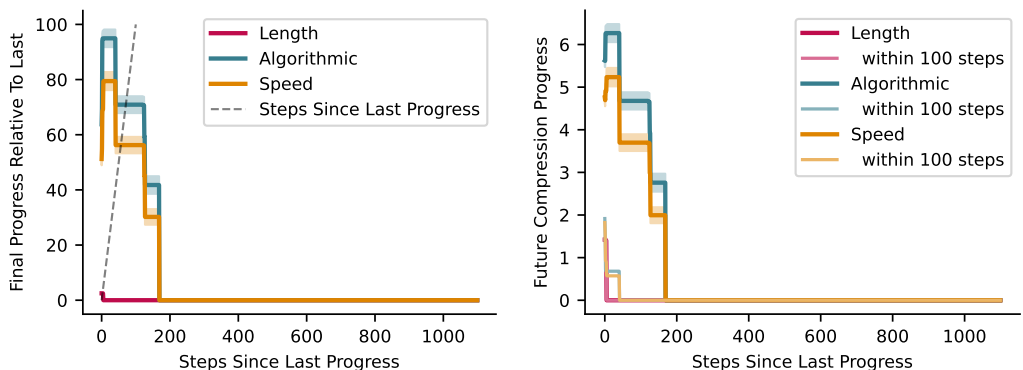


Figure 6: BF. Analogous to Figures 4 and 5. In all three figures, results are plotted with the standard error. It is noticeable only here.

6 CONCLUSION & FUTURE WORK

We have argued that *interestingness* is a necessary algorithmic heuristic for autonomous open-ended intelligence. By formalizing interestingness as a prospective assessment of future compression progress, we have grounded the concept in the rigorous framework of Algorithmic Information Theory. Our analysis of priors over computable objects—Length, Algorithmic, and Speed—reveals that the “inductive property of interestingness” is theoretically and empirically viable: past compression progress can indeed signal the potential for future discovery. However, this signal is highly sensitive to the nature of the underlying prior and the recency of the progress.

Especially this reliance on specific priors warrants further investigation. An inversion of our inquiry could lead to further insights: if we take the existence of inductive interestingness—the property that past learning progress reliably predicts future insight—as a given, what must be true about the underlying distribution of objects? This is a fundamental question for recursive self-improvement: how must a “world” be structured to enable never-ending progress? Our previous analysis of universal priors suggests a somewhat limited memory; the probability of future progress was largely determined by the recency of the last drop. However, in structured domains, we intuitively expect that the *entire history* of progress matters. We do not expect long, steady sequences of discovery to terminate abruptly. This property of sustained, multi-level learnability can be characterized as

scale-free (Barabási & Albert, 1999) emergence. In algorithmic terms, drops in the P_x profile represent transitions between different levels of description or emergence (Bédard & Bergeron, 2022). If an object exhibits scale-free emergence, its profile contains a steady continuation of such drops across many orders of magnitude of complexity. Natural phenomena, such as biological systems or weather and climate, appear to possess this property: they provide an almost endless well of regularities where each discovery uncovers a new layer of puzzles. Also synthetic artifacts with similar characteristics exist: for example certain fractals like the Mandelbrot set, or cellular automata like Conway’s Game of Life. For such objects, we would expect via the $P_x \leftrightarrow D_x$ correspondence that past compression progress is a robust indicator of future learnability. While Jansma & Hoel (2025) have shown that such objects exhibiting causal contributions which are spread out across many levels of coarse-graining can be engineered, it remains a challenge to formally define the conditions required for a prior to prefer these scale-free structures.

Furthermore, we believe that moving beyond human-in-the-loop systems and towards truly self-improving open-ended intelligence, *introspective assessment* of our models is necessary. Current machine learning objectives focus almost exclusively on minimizing post-hoc loss. To achieve autonomy, we must develop introspective models capable of predicting their own learning progress. This requires architectures that do not just compress data, but explicitly model their own “compression frontier”—identifying where an increase in computational effort is likely to yield the highest gain. In the context of modern LLMs, this frontier is increasingly defined by test-time compute and Chain-of-Thought reasoning steps. An autonomous agent must have the ability to decide whether an object is boring (meaning additional runtime will not yield further compression) or interesting (meaning it holds the potential for a complexity drop given more reasoning tokens). Crucially, the exact nature of this computational effort requires a more rigorous taxonomy. Whether defined by longer recurrence, increased thinking steps, additional gradient descent iterations, or expanded model capacity, the underlying algorithmic relationship between these diverse resources remains a vital open question for future research.

By shifting our focus from pure learning to the principled selection of *what* to learn, we can begin to build systems that do not merely solve the tasks we give them, but autonomously seek to discover the richness of the universe they inhabit.

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A ADDITIONAL DISCUSSION

A.1 OPEN-ENDEDNESS EMERGES FROM SOCIAL MULTI-AGENT INTERACTIONS

A compelling alternative is that interestingness is not an intrinsic property of objects, but an emergent phenomenon of social interaction. Such settings span from competitive two-player games (Schmidhuber, 1997) to complex multi-agent ecosystems involving theory-of-mind and embedded agency (Meulemans et al., 2025). There, learning potential is often quantified through disagreement: if a population of models exhibits high variance in their predictions of an object, that object is deemed to contain unresolved, learnable structure (Pathak et al., 2019; Shyam et al., 2019; Sekar et al., 2020; Sancaktar et al., 2022). Similarly, the notion of impressiveness (Lehman & Stanley, 2012) suggests that an artifact is interesting if it serves as a “proof of work”—a hard-to-reach state that signals computational effort to other agents. While we recognize the power of these social dynamics, we argue they do not bypass the fundamental problem of interestingness, they merely decentralize it. Whether an agent is filtering synthetic data or evaluating a peer’s “impressive” achievement, it is still performing a prospective assessment of future progress. Even a co-evolutionary ecosystem (Pugh et al., 2016; Brant & Stanley, 2017) can be viewed as a single macro-agent in a self-generating loop. Therefore, we believe that the inductive properties of learnability we analyze in the following sections remain relevant to these social systems.

A.2 WE CAN PREDICT HUMAN-DISTILLED INTERESTINGNESS VIA LLMs

A compelling alternative to explicit algorithmic measures is the use of LLMs as a “Model of Interestingness”. The OMNI framework (Zhang et al., 2023; Faldor et al., 2024) argues that because LLMs are trained on large amounts of human-generated data, they have already internalized nuanced human notions of what is worthwhile and novel. By prompting an LLM to generate tasks that a human would find interesting, these systems can successfully navigate near-infinite task spaces while avoiding uninspiring or repetitive data.

However, we must consider the long-term limits of this approach. OMNI fundamentally emulates human capacity for nuanced judgment; it effectively asks: “What would a human want to learn next?” While LLMs possess rich commonsense priors, what yields learning progress for a human is not necessarily identical to what yields progress for an artificial subject. Humans and our current AI systems possess vastly different learning architectures and existing knowledge bases. Perhaps, eventually we will converge to a shared conception of interestingness among general intelligences—whether artificial or human. But if the goal is to reach general intelligence through an open-ended process, we cannot assume that the current state of LLMs is already sufficiently close. Furthermore, LLMs currently have no privileged insight into their own internal compression frontiers—they are effectively selecting data based on an external social heuristic rather than an introspective assessment of their own potential for insight. While it is conceivable that a model eventually bootstraps a sophisticated self-understanding, true autonomous open-endedness likely requires the ability to identify potential progress in domains that human intuition hasn’t yet charted.

A.3 INDUCTIVE INTERESTINGNESS IS LIMITED, WE SHOULD CARE ABOUT CONTENT OVER CURVES

A valid critique of the inductive property of interestingness is its apparent disregard for *content*. One might argue that a system should not predict future progress based on the shape of a complexity profile, but rather on the semantic nature of the object itself. From this perspective, an agent can recognize a task as interesting because it identifies familiar motifs—physics-like patterns, linguistic structures, or causal hierarchies—which tend to yield to further analysis.

We can make the argument, however, that accounting for the content is analogous to choosing what kind of prior to use. To an agent with no domain knowledge, the curve is the only universal signal available. But as a system matures, it develops “meta-compression” models—essentially domain-specific priors. These models allow the agent to classify strings into different “structural families.” For instance, a string known to contain large amounts of random noise can prompt a prior algorithmically favoring objects that plateau early, while a string representing a mathematical proof might elicit a prior that favors deep objects and steady progress over many drops.

B MATHEMATICAL PROOFS

We start with some preliminaries. From Equations (2) and (3), we learn that the number of strings with a profile $O(C(P) + \log n_P)$ -close to the given profile P is

$$2^{k_P - m_P + O(\log n_P)}. \quad (7)$$

Here, we set ϵ to be a small constant value, which also justifies replacing $m_P(\epsilon)$, as defined by Vereshchagin & Shen (2016) for technicalities of their proof, with m_P .

A partial profile \hat{P} characterizes sets containing a certain x , up to—but not including—complexity t . It is described by the values

$$\begin{aligned} \hat{n} &:= \min\{r \mid (0, r) \in \hat{P}\}, \\ \hat{k} &:= \min\{r + s \mid (s, r) \in \hat{P}\}, \text{ and} \\ \hat{m} &:= \min\{r \mid (r, \hat{k} - r) \in \hat{P}\}, \end{aligned}$$

with $0 < \hat{m} < t \leq \hat{k} < \hat{n}$. We want to quantify how many strings there are for any possible continuation of the partial profile \hat{P} beyond t to a full profile $P \supset \hat{P}$ with $\{(i, j) \in P \mid i < t\} = \hat{P}$. For all these profiles P , $n_P = \hat{n}$, since the start of the profile is \hat{P} . For the continuations, we distinguish two different scenarios: Either there is no further drop after \hat{P} , or there is at least one more drop. Let's start with the no-further-drop case, where progress *stops* before t . Here, the values \hat{n} , \hat{k} and \hat{m} characterize also the full profile. This means the number of strings that follow profile \hat{P} with no further drops is, according to Equation 7,

$$N_{\text{stop}} := 2^{\hat{k} - \hat{m} + O(\log \hat{n})}.$$

Defining the constant $g_{\text{stop}} := 2^{O(\log \hat{n})}$, we write this as

$$N_{\text{stop}} = g_{\text{stop}} 2^{\hat{k} - \hat{m}}. \quad (8)$$

The constant g_{stop} has a specific value, associated with the number of strings having a profile with the last drop set by \hat{m} and \hat{k} .

If there is at least one more drop, the number of strings with a certain profile depends on the exact position of the last drop, characterized by k and m . Since in this case, the complexity m of the last drop is at least t , we describe further drop condition as “ $m \geq t$ ”. The total number of strings following profile \hat{P} and then exhibiting a further drop is the sum of strings with a particular drop position, summed over all possible drop positions:

$$N_{\text{drop}} := \sum_{m=t}^{\hat{k}-1} \sum_{k=m}^{\hat{k}-1} 2^{k-m+O(\log \hat{n})}.$$

The outer sum enumerates the position of the drops along the complexity axis, the inner sum the size of the drop along the log-size axis. We again re-write this using a constant $g_{\text{drop}} := \sum_{m=t}^{\hat{k}-1} \sum_{k=m}^{\hat{k}-1} 2^{O(\log \hat{n})} = 2^{O(\log \hat{n})}$:

$$\begin{aligned} N_{\text{drop}} &= \sum_{m=t}^{\hat{k}-1} \sum_{k=m}^{\hat{k}-1} 2^{O(\log \hat{n})} \sum_{m=t}^{\hat{k}-1} \sum_{k=m}^{\hat{k}-1} 2^{k-m} \\ &= g_{\text{drop}} \sum_{m=t}^{\hat{k}-1} \sum_{k=m}^{\hat{k}-1} 2^{k-m}. \end{aligned} \quad (9)$$

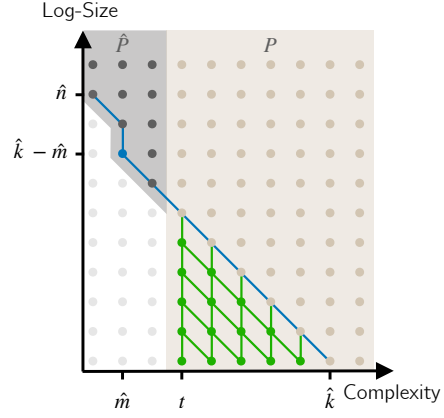


Figure 7: Continuation of profile \hat{P} with no further drop (blue), and all possible continuations with $m \geq t$ (green).

We have to consider a detail: Equation 7 does not quantify the number of strings with a profile of that matches exactly P , but the number of strings with a profile *close* to P . Hence, in the sums above, we may count the same string multiple times. However, the following Lemma shows that this double-counting does not contribute much.

Lemma B.1. *Consider the set of valid profiles with fixed n , and all allowed last drop positions parametrized by (m, k) -tuples. There are at most $O(k_x^2 \log^2 n)$ profiles from this set to which the profile P_x of a given string x , with $n_x = n$, is $O(C(P) + O(\log n))$ -close.*

Proof. Let (m_x, k_x) be the parameters characterizing the last drop of x 's profile P_x . To be $O(C(P) + O(\log n))$ -close to P_x , the last drop² of any profile P has to occur within a square with a size of exactly that precision around the point $(m_x, k_x - m_x)$. There are only $O(C(P)^2 + O(\log n)^2)$ points in such a square. From the fact that P can always be represented as a list of integers not bigger than n , it follows that $C(P) \leq \hat{k} \log n$. Therefore, the square of acceptable positions of the last drop contains $O(k_x^2 \log^2 n)$ points, each with their own valid profile. \square

The double-counting can thus be absorbed into the polynomial constant g_{drop} . Also this constant has a specific value, depending on the precise number of strings with each of the different last drops, and the mentioned double counting. We keep the explicit constants because they can be reduced: For example, $\frac{g_{\text{drop}}}{g_{\text{drop}}} = 1$, whereas in general $\frac{2^{O(\log \hat{n})}}{2^{O(\log \hat{n})}} = 2^{O(\log \hat{n}) - O(\log \hat{n})} = 2^{O(\log \hat{n})}$.

Each point in a profile corresponds to a two-part description of a string. The priors introduced in Section 4 give probabilistic weights to strings, depending on these descriptions.

B.1 LENGTH PRIOR

For the Length Prior L (Equation 4), the probability of a string depends only on its length, i.e. on the value \hat{n} of its profile P (no matter if partial or non-partial, since the length is always determined from the start). Since all strings we are considering share the same initial partial profile \hat{P} , they have the same probability, and we can base our following calculations purely on the relative number of strings.

B.1.1 PROOF OF PROPOSITION 4.1(I): PROBABILITY OF A FURTHER DROP

Proof. The probability that after \hat{P} there will be a further drop is simply the fraction $\frac{N_{\text{drop}}}{N_{\text{drop}} + N_{\text{stop}}}$ of strings exhibiting a further drop. N_{drop} (Equation 9) can be simplified:

$$\begin{aligned} N_{\text{drop}} &= g_{\text{drop}} \sum_{m=t}^{\hat{k}-1} \sum_{k=m}^{\hat{k}-1} 2^{k-m} \\ &= g_{\text{drop}} \left(2^{\hat{k}-t+1} - \hat{k} + t - 2 \right) \\ &= 2^{O(\log \hat{n})} 2^{\hat{k}-t} \end{aligned} \tag{10}$$

For $\hat{k} - t \gg 0$, we can approximate

$$N_{\text{drop}} \approx g_{\text{drop}} 2^{\hat{k}-t+1}. \tag{11}$$

²The drop has to be of size $\Omega(C(P) + \log n)$. We are always assuming drop sizes that are more than logarithmic.

With that, we can calculate the probability of a further drop:

$$\begin{aligned}
p_L(m \geq t \mid \hat{P}) &= \frac{N_{\text{drop}}}{N_{\text{drop}} + N_{\text{stop}}} \\
&= \frac{1}{1 + \frac{N_{\text{stop}}}{N_{\text{drop}}}} \\
&= \frac{1}{1 + \frac{2^{O(\log \hat{n})} 2^{\hat{k} - \hat{m}}}{2^{O(\log \hat{n})} 2^{\hat{k} - t}}} \\
&= 2^{-(t - \hat{m}) + O(\log \hat{n})}
\end{aligned} \tag{12}$$

It diminishes exponentially as the gap between the complexity of the last observed drop \hat{m} and the observation cutoff t grows. \square

B.1.2 PROBABILITY OF SPECIFIC DROP POSITION, GIVEN $m \geq t$

Given \hat{P} and $m \geq t$, the number of strings with a profile where the last drop occurs at a specific complexity m^* is

$$\begin{aligned}
&\sum_{m=t}^{\hat{k}-1} \sum_{k=m}^{\hat{k}-1} \mathbf{1}[m = m^*] 2^{k-m+O(\log \hat{n})} \\
&= \sum_{k=m^*}^{\hat{k}-1} 2^{k-m^*+O(\log \hat{n})} \\
&= 2^{\hat{k}-m^*+O(\log \hat{n})}.
\end{aligned} \tag{13}$$

The probability of this drop position is hence

$$\begin{aligned}
p_L(m = m^* \mid \hat{P}, m \geq t) &= \frac{2^{\hat{k}-m^*+O(\log \hat{n})}}{N_{\text{drop}}} \\
&= \frac{2^{\hat{k}-m^*+O(\log \hat{n})}}{2^{\hat{k}-t+O(\log \hat{n})}} \\
&= 2^{t-m^*+O(\log \hat{n})}.
\end{aligned} \tag{14}$$

This shows that the probability of the last drop occurring at m^* decays exponentially as m^* moves away from the current threshold t . We now perform a parallel derivation for the final complexity k^* . The number of strings with a specific complexity k^* , given \hat{P} and $m \geq t$, is

$$\begin{aligned}
\sum_{m=t}^{\hat{k}-1} \sum_{k=m}^{\hat{k}-1} \mathbf{1}[k = k^*] 2^{k-m+O(\log \hat{n})} &= \sum_{m=t}^{k^*} 2^{k^*-m+O(\log \hat{n})} \\
&= 2^{k^*-t+O(\log \hat{n})},
\end{aligned} \tag{15}$$

and the probability of this complexity is

$$\begin{aligned}
p_L(k = k^* \mid \hat{P}, m \geq t) &= \frac{2^{k^*-t+O(\log \hat{n})}}{N_{\text{drop}}} \\
&= \frac{2^{k^*-t+O(\log \hat{n})}}{2^{\hat{k}-t+O(\log \hat{n})}} \\
&= 2^{k^*-\hat{k}+O(\log \hat{n})}.
\end{aligned} \tag{16}$$

B.1.3 PROOF OF PROPOSITION 4.1(II): EXPECTED VALUES, GIVEN $m \geq t$

Proof. Now we calculate the expected values of m and k under the Length Prior, under the condition of a further drop $m \geq t$. For that, we first compute the sums of the number of strings with a certain profile, multiplied with the respective value of m and k :

$$\begin{aligned} M_L &:= g_{\text{drop}} \sum_{m=t}^{\hat{k}-1} \sum_{k=m}^{\hat{k}-1} m 2^{k-m} \\ &= g_{\text{drop}} \left(2^{\hat{k}-t+1} (t+1) - 2\hat{k} + \frac{1}{2} \left(-\hat{k}^2 - 3\hat{k} + t^2 - t - 4 \right) \right), \end{aligned} \quad (17)$$

and

$$\begin{aligned} K_L &:= g_{\text{drop}} \sum_{m=t}^{\hat{k}-1} \sum_{k=m}^{\hat{k}-1} k 2^{k-m} \\ &= g_{\text{drop}} \left(2^{\hat{k}-t+1} (\hat{k} - 2) + \frac{1}{2} \left(-\hat{k}^2 + \hat{k} + t^2 - 5t + 8 \right) \right) \end{aligned} \quad (18)$$

For $\hat{k} - t \gg 0$, these values simplify to

$$M_L \approx g_{\text{drop}} 2^{\hat{k}-t+1} (t+1), \quad (19)$$

and

$$K_L \approx g_{\text{drop}} 2^{\hat{k}-t+1} (\hat{k} - 2). \quad (20)$$

The expected values are then, using Equation 11,

$$\begin{aligned} \mathbb{E}_{x \sim L}[m \mid \hat{P}, m \geq t] &= \frac{M_L}{N_{\text{drop}}} \\ &\approx \frac{g_{\text{drop}} 2^{\hat{k}-t+1} (t+1)}{g_{\text{drop}} 2^{\hat{k}-t+1}} \\ &= t+1, \end{aligned} \quad (21)$$

and

$$\begin{aligned} \mathbb{E}_{x \sim L}[k \mid \hat{P}, m \geq t] &= \frac{K_L}{N_{\text{drop}}} \\ &\approx \frac{g_{\text{drop}} 2^{\hat{k}-t+1} (\hat{k} - 2)}{g_{\text{drop}} 2^{\hat{k}-t+1}} \\ &= \hat{k} - 2. \end{aligned} \quad (22)$$

Under the Length Prior, further drops are most likely to occur close to the observation threshold t , and to be small—leading to a high final complexity. \square

B.1.4 PROOF OF PROPOSITION 4.1(III): EXPECTED PROGRESS, CONDITIONED ONLY ON \hat{P}

Proof. Observing the partial profile \hat{P} , the difference in complexity between the last observed drop at \hat{m} , and the expected last drop is

$$\begin{aligned} \mathbb{E}_{x \sim L}[m \mid \hat{P}] - \hat{m} &= \hat{m} \left(1 - p_L(m \geq t \mid \hat{P}) \right) + \mathbb{E}_{x \sim L}[m \mid \hat{P}, m \geq t] p_L(m \geq t \mid \hat{P}) - \hat{m} \\ &= p_L(m \geq t \mid \hat{P}) \left(\mathbb{E}_{x \sim L}[m \mid \hat{P}, m \geq t] - \hat{m} \right). \end{aligned} \quad (23)$$

Since the second factor (the complexity difference) is at most linear in \hat{k} , it is dominated by the $2^{O(\log \hat{n})}$ of the probability, and thus, with Equations 12 and 21,

$$\mathbb{E}_{x \sim L}[m \mid \hat{P}] - \hat{m} = 2^{-(t-\hat{m})+O(\log \hat{n})}. \quad (24)$$

The expected compression progress, relative to \hat{k} , is

$$\begin{aligned} \hat{k} - \mathbb{E}_{x \sim L}[k \mid \hat{P}] &= \hat{k} - \hat{k} \left(1 - p_L(m \geq t \mid \hat{P})\right) - \mathbb{E}_{x \sim L}[k \mid \hat{P}, m \geq t] p_L(m \geq t \mid \hat{P}) \\ &= p_L(m \geq t \mid \hat{P}) \left(\hat{k} - \mathbb{E}_{x \sim L}[k \mid \hat{P}, m \geq t]\right), \end{aligned} \quad (25)$$

and again, with Equations 12 and 22,

$$\hat{k} - \mathbb{E}_{x \sim L}[k \mid \hat{P}] = 2^{-(t-\hat{m})+O(\log \hat{n})}. \quad (26)$$

□

B.1.5 PROBABILITY OF A DROP WITHIN A GIVEN WINDOW

We do not only care about the total future compression potential of a string, but perhaps even more about the near-term potential. To study this, we define the interval $d := \{t, \dots, t + \Delta c - 1\}$, demarcating a specific window of interest (the green shaded area in Figure 8). We want to establish the probability of a drop happening within d , and the expected compression progress within the interval.

Since so far, we only cared about the indefinite future, we were able to only rely on the properties associated with the *last* drop. But now we care about *any* drop within d . This requires some additional care.

Apart from the case where the observed drop defined by \hat{m} and \hat{k} remains the last one, we distinguish three more scenarios:

- **Drop profiles (Green):** The last drop occurs within the window d .
- **Void profiles (Red):** The last drop occurs after $t + \Delta c$, but the drop is too shallow to allow any intermediate drops within d .
- **Perhaps profiles (Yellow):** The last drop occurs after the window, but the complexity gap is large enough that a preceding drop could exist within d .

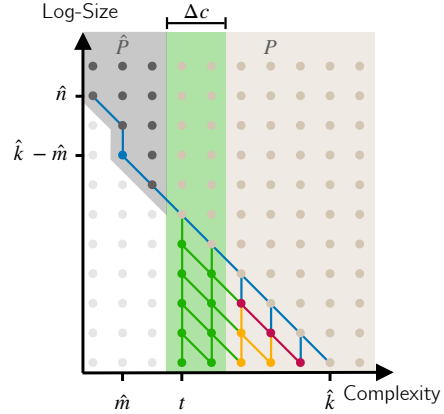


Figure 8: Continuation of profile \hat{P} . In blue, continuation with no further drop. In green, continuations with their last drop within Δc after t . In red, continuations that have their last drop after the interval, and cannot have any drop within the interval. In yellow, continuations that have their last drop after the interval, but may have another drop within the interval.

The number of strings in the “drop” is, defining $g(d)_{\text{drop}} := \sum_{m=t}^{t+\Delta c-1} \sum_{k=m}^{\hat{k}-1} 2^{O(\log \hat{n})} = 2^{O(\log \hat{n})}$

$$\begin{aligned} N(d)_{\text{drop}} &:= g(d)_{\text{drop}} \sum_{m=t}^{t+\Delta c-1} \sum_{k=m}^{\hat{k}-1} 2^{k-m} \\ &= g(d)_{\text{drop}} \left(2^{\hat{k}-t+1} - 2^{\hat{k}-t-\Delta c+1} - \Delta c\right) \\ &= 2^{\hat{k}-t+O(\log \hat{n})} - 2^{\hat{k}-t-\Delta c+O(\log \hat{n})}, \end{aligned} \quad (27)$$

which for $\hat{k} - t \gg 0$ can be simplified to

$$N(d)_{\text{drop}} \approx g(d)_{\text{drop}} \left(1 - 2^{-\Delta c}\right) 2^{\hat{k}-t}. \quad (28)$$

Defining $g(d)_{\text{void}}$ and $g(d)_{\text{perhaps}}$ equivalently, we get

$$\begin{aligned}
N(d)_{\text{void}} &:= g_{\text{void}}(d) \sum_{m=t+\Delta c}^{\hat{k}-1} 2^{\hat{k}-1-m} \\
&= g_{\text{void}}(d) \left(2^{\hat{k}-t-\Delta c} - 1 \right) \\
&= 2^{\hat{k}-t-\Delta c+O(\log \hat{n})}
\end{aligned} \tag{29}$$

and

$$\begin{aligned}
N(d)_{\text{perhaps}} &:= g(d)_{\text{perhaps}} \sum_{m=t+\Delta c}^{\hat{k}-2} \sum_{k=m}^{\hat{k}-2} 2^{k-m} \\
&= g(d)_{\text{perhaps}} \left(2^{\hat{k}-t-\Delta c} + \Delta c - \hat{k} + t - 1 \right) \\
&= 2^{\hat{k}-t-\Delta c+O(\log \hat{n})}.
\end{aligned} \tag{30}$$

Recall the value N_{drop} defined in Equation 9, that quantifies all strings with a drop after \hat{P} . The number of strings with a drop within d is lower bounded by $N_{\text{drop}} - (N(d)_{\text{void}} + N(d)_{\text{perhaps}})$ and upper bounded by $N_{\text{drop}} - N(d)_{\text{void}}$. This allows us to bound the probability:

$$\frac{N_{\text{drop}} - (N(d)_{\text{void}} + N(d)_{\text{perhaps}})}{N_{\text{drop}} + N_{\text{stop}}} \leq p_L(\text{drop within } d \mid \hat{P}) \leq \frac{N_{\text{drop}} - N(d)_{\text{void}}}{N_{\text{drop}} + N_{\text{stop}}} \tag{31}$$

Notice that, due to the $O(\log \hat{n})$ -term in the exponent,

$$N(d)_{\text{void}} + N(d)_{\text{perhaps}} = 2^{\hat{k}-t-\Delta c+O(\log \hat{n})} = N(d)_{\text{void}}. \tag{32}$$

Hence,

$$\begin{aligned}
p_L(\text{drop within } d \mid \hat{P}) &= \frac{N_{\text{drop}} - N(d)_{\text{void}}}{N_{\text{drop}} + N_{\text{stop}}} \\
&= \frac{2^{O(\log \hat{n})} 2^{\hat{k}-t} - 2^{O(\log \hat{n})} 2^{\hat{k}-t-\Delta c}}{2^{O(\log \hat{n})} 2^{\hat{k}-t} + 2^{O(\log \hat{n})} 2^{\hat{k}-\hat{m}}} \\
&= 2^{\hat{m}-t+O(\log \hat{n})} (1 - 2^{-\Delta c}).
\end{aligned} \tag{33}$$

We can see that the probability of a further drop is highly concentrated close to t , and intervals of larger size Δc increase the probability by an exponentially vanishing amount. Furthermore, since the probability of the last drop occurring within d is $p_L(\text{last drop within } d \mid \hat{P}) = \frac{N_{\text{drop}} - (N(d)_{\text{void}} + N(d)_{\text{perhaps}})}{N_{\text{drop}} + N_{\text{stop}}}$, it follows that, up to the usual logarithmic precision in the exponents,

$$p_L(\text{last drop within } d \mid \hat{P}) = p_L(\text{drop within } d \mid \hat{P}). \tag{34}$$

This means that the profiles with the last drop within d dominate the probability of having *any* drop within d .

B.1.6 EXPECTED PROGRESS WITHIN A GIVEN WINDOW

Similarly to Sections B.1.3 and B.1.4, we can calculate the expected compression progress within d . Again, we start with the expected value for k , given that the last drop is within the interval d . The weighted number of corresponding strings is

$$\begin{aligned}
K_L(d) &:= g(d)_{\text{drop}} \sum_{m=t}^{t+\Delta c-1} \sum_{k=m}^{\hat{k}-1} k 2^{k-m} \\
&= g(d)_{\text{drop}} (\hat{k} - 2) (2^{\hat{k}-t+1} - 2^{\hat{k}-t-\Delta c+1}) + g(d)_{\text{drop}} \frac{\Delta c}{2} (5 - 2t - \Delta c),
\end{aligned} \tag{35}$$

$$K_L(d) \approx g(d)_{\text{drop}} (\hat{k} - 2) (1 - 2^{-\Delta c}) 2^{\hat{k}-t}. \tag{36}$$

With that, the expected value for k , assuming that the last drop happens within d and using Equation 28, is

$$\mathbb{E}_{x \sim L}[k \mid \hat{P}, \text{ last drop within } d] = \frac{K_L(d)}{N(d)_{\text{drop}}} \approx \hat{k} - 2. \quad (37)$$

From Equation 34 it follows that also

$$\mathbb{E}_{x \sim L}[k \mid \hat{P}, \text{ drop within } d] \approx \hat{k} - 2.$$

Now we can compute the expected compression progress within d . Let $k_{<t+\Delta c}$ be the complexity of the last drop before $t + \Delta c$ (but not necessarily the last drop of the whole profile). That means, $k_{<t+\Delta c}$ represents the “achieved complexity” within the window, regardless of whether the string eventually becomes even simpler later in the profile. Then,

$$\begin{aligned} \hat{k} - \mathbb{E}_{x \sim L}[k_{<t+\Delta c} \mid \hat{P}] &= p_L(\text{drop within } d \mid \hat{P}) \left(\hat{k} - \mathbb{E}_{x \sim L}[k \mid \hat{P}, \text{ drop within } d] \right) \\ &\approx 2^{\hat{m}-t+O(\log \hat{n})} (1 - 2^{-\Delta c}). \end{aligned} \quad (38)$$

B.2 ALGORITHMIC PRIOR

The Length Prior from the previous section has the property that all strings of length \hat{n} have the same probability. This is not the case for the Algorithmic Prior M (Equation 5): Here, the probability of a string is determined by its complexity, i.e. the value k of its profile P . Nevertheless, we can proceed in a manner similar to before.

B.2.1 PROOF OF PROPOSITION 4.2(1): PROBABILITY OF FURTHER DROP

Proof. Instead of simply counting strings, we weigh each string with its prior probability. The unnormalized statistical weight of profiles with no further drop under the Algorithmic Prior M is

$$\begin{aligned} Z_{\text{stop}}^M &:= 2^{\hat{k}-\hat{m}+O(\log \hat{n})} 2^{-\hat{k}} \\ &= 2^{O(\log \hat{n})} 2^{-\hat{m}} \\ &= g_{\text{stop}} 2^{-\hat{m}}. \end{aligned} \quad (39)$$

With at least one more drop, the weighted number of strings is

$$\begin{aligned} Z_{\text{drop}}^M &:= \sum_{m=t}^{\hat{k}-1} \sum_{k=m}^{\hat{k}-1} 2^{k-m+O(\log \hat{n})} 2^{-k} \\ &= g_{\text{drop}} \sum_{m=t}^{\hat{k}-1} \sum_{k=m}^{\hat{k}-1} 2^{-m} \\ &= g_{\text{drop}} \left((\hat{k} - t - 1) 2^{1-t} + 2^{1-\hat{k}} \right) \\ &= 2^{-t+O(\log \hat{n})} + 2^{-\hat{k}+O(\log \hat{n})} \end{aligned} \quad (40)$$

For $\hat{k} \gg t$, this is

$$Z_{\text{drop}}^M \approx g_{\text{drop}} (\hat{k} - t - 1) 2^{1-t}. \quad (41)$$

This yields the probability of a further drop

$$\begin{aligned}
p_M(m \geq t \mid \hat{P}) &= \frac{1}{1 + \frac{Z_{\text{stop}}^M}{Z_{\text{drop}}^M}} \\
&= \frac{1}{1 + \frac{2^{O(\log \hat{n})}}{2^{O(\log \hat{n})}} \frac{2^{-\hat{m}}}{\left((\hat{k}-t-1)2^{1-t} + 2^{1-\hat{k}} \right)}} \\
&= \frac{1}{1 + \frac{2^{-\hat{m}+O(\log \hat{n})}}{2^{-t+O(\log \hat{n})} + 2^{-\hat{k}+O(\log \hat{n})}}}. \tag{42}
\end{aligned}$$

Here, we have a dependence on \hat{k} in the denominator, which does not exist for the Length Prior (compare Equation 12). However, this dependence vanishes for $\hat{k} - t \gg 0$:

$$\begin{aligned}
p_M(m \geq t \mid \hat{P}) &= \frac{1}{1 + \frac{Z_{\text{stop}}^M}{Z_{\text{drop}}^M}} \\
&\approx \frac{1}{1 + \frac{2^{-\hat{m}+O(\log \hat{n})}}{2^{-t+O(\log \hat{n})}}} \\
&= 2^{-(t-\hat{m})+O(\log \hat{n})} \tag{43}
\end{aligned}$$

□

B.2.2 PROBABILITY OF SPECIFIC DROP POSITION, GIVEN $m \geq t$

Given \hat{P} and $m \geq t$, the number of strings with a profile where the last drop occurs at a specific complexity m^* , weighted by the Algorithmic Prior, is:

$$\begin{aligned}
&\sum_{m=t}^{\hat{k}-1} \sum_{k=m}^{\hat{k}-1} \mathbf{1}[m = m^*] 2^{k-m+O(\log \hat{n})} 2^{-k} \\
&= \sum_{k=m^*}^{\hat{k}-1} 2^{-m^*+O(\log \hat{n})} \\
&= 2^{-m^*+O(\log \hat{n})}. \tag{44}
\end{aligned}$$

The probability of this drop position is hence:

$$\begin{aligned}
p_M(m = m^* \mid \hat{P}, m \geq t) &= \frac{2^{-m^*+O(\log \hat{n})}}{Z_{\text{drop}}^M} \\
&= \frac{2^{-m^*+O(\log \hat{n})}}{2^{-t+O(\log \hat{n})} + 2^{-\hat{k}+O(\log \hat{n})}}. \tag{45}
\end{aligned}$$

For $\hat{k} - t \gg 0$, this simplifies to

$$p_M(m = m^* \mid \hat{P}, m \geq t) \approx 2^{t-m^*+O(\log \hat{n})}, \tag{46}$$

which matches the Length Prior result in Equation 14.

In contrast, the weighted number of strings with a specific complexity k^* , given \hat{P} and $m \geq t$, is:

$$\begin{aligned}
&2^{O(\log \hat{n})} \sum_{m=t}^{\hat{k}-1} \sum_{k=m}^{\hat{k}-1} \mathbf{1}[k = k^*] 2^{k-m} 2^{-k} \\
&= 2^{O(\log \hat{n})} \sum_{m=t}^{k^*} 2^{-m} \\
&= 2^{-t+O(\log \hat{n})} - 2^{-k^*+O(\log \hat{n})}. \tag{47}
\end{aligned}$$

The probability of this complexity is:

$$p_M(k = k^* \mid \hat{P}, m \geq t) = \frac{2^{-t+O(\log \hat{n})} - 2^{-k^*+O(\log \hat{n})}}{2^{-t+O(\log \hat{n})} + 2^{-\hat{k}+O(\log \hat{n})}}. \quad (48)$$

For $\hat{k} - t \gg 0$, this yields:

$$p_M(k = k^* \mid \hat{P}, m \geq t) \approx 1 - 2^{t-k^*+O(\log \hat{n})}. \quad (49)$$

Unlike the Length Prior (Equation 16), where the distribution of k vanishes exponentially as it moves away from \hat{k} , the Algorithmic Prior results in a distribution that is nearly uniform for $k^* \gg t$.

B.2.3 PROOF OF PROPOSITION 4.2(II): EXPECTED VALUES, GIVEN $m \geq t$

Probabilistically weighted sums of the number of strings with a certain profile, multiplied with the respective value of m and k :

$$\begin{aligned} M_M &:= g_{\text{drop}} \sum_{m=t}^{\hat{k}-1} \sum_{k=m}^{\hat{k}-1} m 2^{-m} \\ &= g_{\text{drop}} \left(\left((\hat{k} - t - 1)(2t + 2) - 4 \right) 2^{-t} + (2\hat{k} + 6) 2^{-\hat{k}} \right) \end{aligned} \quad (50)$$

$$\begin{aligned} K_M &:= g_{\text{drop}} \sum_{m=t}^{\hat{k}-1} \sum_{k=m}^{\hat{k}-1} k 2^{-m} \\ &= g_{\text{drop}} \left(\left((\hat{k} - t - 1)(\hat{k} + t) - 2 \right) 2^{-t} + (2\hat{k} + 2) 2^{-\hat{k}} \right) \end{aligned} \quad (51)$$

For $\hat{k} - t \gg 0$, these values simplify to

$$M_M \approx g_{\text{drop}} (\hat{k} - t - 1)(2t + 2) 2^{-t}, \quad (52)$$

$$K_M \approx g_{\text{drop}} (\hat{k} - t - 1)(\hat{k} + t) 2^{-t}. \quad (53)$$

The expected values are then, using Equation 41

$$\begin{aligned} \mathbb{E}_{x \sim M}[m \mid \hat{P}, m \geq t] &= \frac{M_M}{Z_{\text{drop}}^M} \\ &\approx \frac{g_{\text{drop}} (\hat{k} - t - 1)(2t + 2) 2^{-t}}{g_{\text{drop}} (\hat{k} - t - 1) 2^{1-t}} \\ &= t + 1, \end{aligned} \quad (54)$$

$$\begin{aligned} \mathbb{E}_{x \sim M}[k \mid \hat{P}, m \geq t] &= \frac{K_M}{Z_{\text{drop}}^M} \\ &\approx \frac{g_{\text{drop}} (\hat{k} - t - 1)(\hat{k} + t) 2^{-t}}{g_{\text{drop}} (\hat{k} - t - 1) 2^{1-t}} \\ &= \frac{\hat{k} + t}{2}. \end{aligned} \quad (55)$$

Here, the approximate uniformity from Equation 49 comes to bear: The expectation for k is significantly lower compared to the Length Prior (compare Equation 22).

B.2.4 PROOF OF PROPOSITION 4.2 (III): EXPECTED PROGRESS, CONDITIONED ONLY ON \hat{P}

Proof. Nevertheless, just as with the Length Prior, the expected progress, conditioned only on \hat{P} is dominated by the drop probability:

$$\mathbb{E}_{x \sim M}[m \mid \hat{P}] - \hat{m} \approx 2^{-(t-\hat{m})+O(\log \hat{n})}, \quad (56)$$

and

$$\hat{k} - \mathbb{E}_{x \sim M}[k \mid \hat{P}] \approx 2^{-(t-\hat{m})+O(\log \hat{n})}. \quad (57)$$

□

B.2.5 PROBABILITY OF A DROP WITHIN A GIVEN WINDOW

Analogously to the Length Prior (Section B.1.5), we calculate the weight of the three cases:

$$\begin{aligned} Z(d)_{\text{drop}}^M &:= g(d)_{\text{drop}} \sum_{m=t}^{t+\Delta c-1} \sum_{k=m}^{\hat{k}-1} 2^{-m} \\ &= g(d)_{\text{drop}} 2^{-t+1} \left((1-2^{-\Delta c}) (\hat{k}-t-1) + \Delta c 2^{-\Delta c} \right) \\ &= 2^{-t+O(\log \hat{n})} - 2^{-t-\Delta c+O(\log \hat{n})} \end{aligned} \quad (58)$$

$$Z(d)_{\text{drop}}^M \approx g(d)_{\text{drop}} 2^{-t+1} (1-2^{-\Delta c}) (\hat{k}-t) \quad (59)$$

$$\begin{aligned} Z(d)_{\text{void}}^M &:= g_{\text{void}}(d) \sum_{m=t+\Delta c}^{\hat{k}-1} 2^{-m-1} \\ &= g_{\text{void}}(d) \left(2^{-t-\Delta c} - 2^{-\hat{k}} \right) \\ &= 2^{-t-\Delta c+O(\log \hat{n})} - 2^{-\hat{k}+O(\log \hat{n})} \end{aligned} \quad (60)$$

$$\begin{aligned} Z(d)_{\text{perhaps}}^M &:= g(d)_{\text{perhaps}} \sum_{m=t+\Delta c}^{\hat{k}-2} \sum_{k=m}^{\hat{k}-2} 2^{-m} \\ &= g(d)_{\text{perhaps}} \left((\hat{k}-t-\Delta c-2) 2^{-t-\Delta c+1} + 2^{-\hat{k}+2} \right) \\ &= 2^{-t-\Delta c+O(\log \hat{n})} + 2^{-\hat{k}+O(\log \hat{n})} \end{aligned} \quad (61)$$

Using Equation 41, we can bound the probability:

$$\frac{Z_{\text{drop}}^M - \left(Z(d)_{\text{void}}^M + Z(d)_{\text{perhaps}}^M \right)}{Z_{\text{drop}}^M + Z_{\text{stop}}^M} \leq p_M(\text{drop within } d \mid \hat{P}) \leq \frac{Z_{\text{drop}}^M - Z(d)_{\text{void}}^M}{Z_{\text{drop}}^M + Z_{\text{stop}}^M} \quad (62)$$

And again, due to the $O(\log \hat{n})$ -term in the exponent,

$$Z(d)_{\text{void}}^M + Z(d)_{\text{perhaps}}^M = 2^{-t-\Delta c+O(\log \hat{n})} - 2^{-\hat{k}+O(\log \hat{n})} = Z(d)_{\text{void}}^M. \quad (63)$$

Hence,

$$\begin{aligned} p_M(\text{drop within } d \mid \hat{P}) &= \frac{Z_{\text{drop}}^M - Z(d)_{\text{void}}^M}{Z_{\text{drop}}^M + Z_{\text{stop}}^M} \\ &= 2^{O(\log \hat{n})} \frac{2^{-t} + 2^{-\hat{k}} - 2^{-t-\Delta c} + 2^{-\hat{k}}}{2^{-t} + 2^{-\hat{k}} + 2^{-\hat{m}}} \\ &= \frac{2^{-t+O(\log \hat{n})} (1-2^{-\Delta c}) + 2^{-\hat{k}+O(\log \hat{n})}}{2^{O(\log \hat{n})} (2^{-t} + 2^{-\hat{k}} + 2^{-\hat{m}})}, \end{aligned} \quad (64)$$

and for $\hat{k} - t \gg 0$,

$$\begin{aligned} p_M(\text{drop within } d \mid \hat{P}) &\approx \frac{2^{-t+O(\log \hat{n})} (1 - 2^{-\Delta c})}{2^{-t+O(\log \hat{n})} + 2^{-\hat{m}+O(\log \hat{n})}} \\ &= 2^{-(t-\hat{m})+O(\log \hat{n})} (1 - 2^{-\Delta c}). \end{aligned} \quad (65)$$

Also here, larger values of Δc contribute only by an exponentially vanishing amount. It again follows that

$$p_M(\text{last drop within } d \mid \hat{P}) = p_M(\text{drop within } d \mid \hat{P}). \quad (66)$$

B.2.6 EXPECTED PROGRESS WITHIN A GIVEN WINDOW

Same procedure as Section B.1.6: The weighted number of corresponding strings is

$$\begin{aligned} K_M(d) &= g(d)_{\text{drop}} \sum_{m=t}^{t+\Delta c-1} \sum_{k=m}^{\hat{k}-1} k 2^{-m} \\ &= g(d)_{\text{drop}} 2^{-t} \left((1 - 2^{-\Delta c}) (\hat{k} + t)(\hat{k} - t - 1) + 2^{-\Delta c} (\Delta c + 2)(\Delta c + 1) - 2 \right), \\ K_M(d) &\approx g(d)_{\text{drop}} 2^{-t} (1 - 2^{-\Delta c}) (\hat{k} + t)(\hat{k} - t). \end{aligned} \quad (67)$$

The expected value for k , assuming that the last drop happens within d and using Equation 59, is

$$\mathbb{E}_{x \sim M}[k \mid \hat{P}, \text{ last drop within } d] = \frac{K_M(d)}{Z(d)_{\text{drop}}^M} \approx \frac{\hat{k} + t}{2}. \quad (68)$$

From Equation 66 it follows that

$$\mathbb{E}_{x \sim M}[k \mid \hat{P}, \text{ drop within } d] \approx \frac{\hat{k} + t}{2}, \quad (69)$$

and

$$\begin{aligned} \hat{k} - \mathbb{E}_{x \sim M}[k_{<t+\Delta c} \mid \hat{P}] &= p_M(\text{drop within } d \mid \hat{P}) \left(\hat{k} - \mathbb{E}_{x \sim M}[k \mid \hat{P}, \text{ drop within } d] \right) \\ &\approx 2^{-(t-\hat{m})+O(\log \hat{n})} (1 - 2^{-\Delta c}). \end{aligned} \quad (70)$$

B.3 COMPARISON BETWEEN LENGTH AND ALGORITHMIC PRIOR

The analytic forms of the further drop probability (Equations 12 and 43), as well as the expected progress (Equations 26 and 57) are identical for Length and Algorithmic Prior. However, we can show that there is in fact a difference between the two priors.

B.3.1 PROOF OF PROPOSITION 4.3(I): RATIO OF DROP PROBABILITIES

Proof. Using Equations 8, 11, 39, and 41, we can calculate the ratio of the drop probabilities:

$$\begin{aligned}
\frac{p_M(m \geq t \mid \hat{P})}{p_L(m \geq t \mid \hat{P})} &= \frac{Z_{\text{drop}}^M (N_{\text{drop}} + N_{\text{stop}})}{N_{\text{drop}} (Z_{\text{drop}}^M + Z_{\text{stop}}^M)} \\
&= \frac{g_{\text{drop}} (\hat{k} - t - 1) 2^{1-t} (g_{\text{drop}} 2^{\hat{k}-t+1} + g_{\text{stop}} 2^{\hat{k}-\hat{m}})}{g_{\text{drop}} 2^{\hat{k}-t+1} (g_{\text{drop}} (\hat{k} - t - 1) 2^{1-t} + g_{\text{stop}} 2^{-\hat{m}})} \\
&= 1 + \frac{(\hat{k} - t - 2) g_{\text{stop}} 2^{\hat{k}-t-\hat{m}+1}}{g_{\text{drop}} (\hat{k} - t - 1) 2^{\hat{k}-2t+2} + g_{\text{stop}} 2^{\hat{k}-t-\hat{m}+1}} \\
&= 1 + \frac{\hat{k} - t - 2}{1 + \frac{g_{\text{drop}}}{g_{\text{stop}}} (\hat{k} - t - 1) 2^{-(t-\hat{m}-1)}} \\
&= 1 + \frac{\hat{k} - t - 2}{1 + 2^{-(t-\hat{m})+O(\log \hat{n})}} \tag{72}
\end{aligned}$$

For $t - \hat{m} \gg \log \hat{n}$, this simplifies to

$$\frac{p_M(m \geq t \mid \hat{P})}{p_L(m \geq t \mid \hat{P})} \approx \hat{k} - t - 1. \tag{73}$$

This result implies that the future drop probability under the Algorithmic Prior is higher than under the Length Prior by a factor approximately linear in the complexity gap $\hat{k} - t$. \square

B.3.2 PROOF OF PROPOSITION 4.3(II): RATIO OF COMPRESSION PROGRESS

Proof. The ratio of expected compression progress is derived by combining the ratio of discovery probabilities with the ratio of conditional gains:

$$\frac{\hat{k} - \mathbb{E}_{x \sim M}[k \mid \hat{P}]}{\hat{k} - \mathbb{E}_{x \sim L}[k \mid \hat{P}]} = \frac{p_M(m \geq t \mid \hat{P}) (\hat{k} - \mathbb{E}_{x \sim M}[k \mid \hat{P}, m \geq t])}{p_L(m \geq t \mid \hat{P}) (\hat{k} - \mathbb{E}_{x \sim L}[k \mid \hat{P}, m \geq t])}. \tag{74}$$

Assuming $\hat{k} - t \gg 0$ and $t - \hat{m} \gg \log \hat{n}$, we substitute the results from Equations 22, 55, and 73:

$$\begin{aligned}
\frac{\hat{k} - \mathbb{E}_{x \sim M}[k \mid \hat{P}]}{\hat{k} - \mathbb{E}_{x \sim L}[k \mid \hat{P}]} &\approx (\hat{k} - t - 1) \frac{\hat{k} - \frac{1}{2} (\hat{k} + t)}{\hat{k} - (\hat{k} - 2)} \\
&= \frac{1}{4} (\hat{k} - t - 1) (\hat{k} - t). \tag{75}
\end{aligned}$$

Thus, the expected compression progress under the Algorithmic Prior is larger than under the Length Prior by a factor that is approximately quadratic in the remaining complexity gap $\hat{k} - t$. \square

B.4 SPEED PRIOR

The Speed Prior (Equation 6) assesses the probability of a string based on both the length and the log runtime of programs computing that string. Using the $P_x \leftrightarrow D_x$ correspondence (Equation 1), we can convert the value m into a (extremely fast growing) runtime, which is punished by the Speed Prior. For any sufficiently large value i , $\log \text{BB}(i) / \log \text{BB}(i+1) \approx 0$.

B.4.1 PROOF OF PROPOSITION 4.4(I): PROBABILITY OF FURTHER DROP

Proof. Unnormalized statistical weight of profiles with no further drop under the Speed Prior S :

$$\begin{aligned} Z_{\text{stop}}^S &:= 2^{\hat{k}-\hat{m}+O(\log \hat{n})} 2^{-\hat{k}-\log \text{BB}(\hat{m})} \\ &= g_{\text{stop}} 2^{-\hat{m}-\log \text{BB}(\hat{m})} \end{aligned} \quad (76)$$

For continuations with at least one more drop, the weight is:

$$\begin{aligned} Z_{\text{drop}}^S &:= \sum_{m=t}^{\hat{k}-1} \sum_{k=m}^{\hat{k}-1} 2^{k-m+O(\log \hat{n})} 2^{-k-\log \text{BB}(m)} \\ &\approx g_{\text{drop}} \sum_{k=t}^{\hat{k}-1} 2^{-t-\log \text{BB}(t)} \\ &= g_{\text{drop}} (\hat{k}-t) 2^{-t-\log \text{BB}(t)}, \end{aligned} \quad (77)$$

where the approximation follows from the fact that the $m = t$ term dominates due to the extreme growth of BB. This yields the probability of a further drop

$$\begin{aligned} p_S(m \geq t \mid \hat{P}) &= \frac{Z_{\text{drop}}^S}{Z_{\text{drop}}^S + Z_{\text{stop}}^S} \\ &\approx \frac{g_{\text{drop}} (\hat{k}-t) 2^{-t-\log \text{BB}(t)}}{g_{\text{drop}} (\hat{k}-t) 2^{-t-\log \text{BB}(t)} + g_{\text{stop}} 2^{-\hat{m}-\log \text{BB}(\hat{m})}} \\ &\approx 0, \end{aligned} \quad (78)$$

since $t > \hat{m}$. □

B.4.2 PROBABILITIES OF SPECIFIC DROP POSITIONS GIVEN A FURTHER DROP

Given \hat{P} and $m \geq t$, we evaluate the weighted number of strings where the last drop occurs at complexity m^* :

$$\begin{aligned} &\sum_{m=t}^{\hat{k}-1} \sum_{k=m}^{\hat{k}-1} \mathbf{1}[m = m^*] 2^{k-m+O(\log \hat{n})} 2^{-k-\log \text{BB}(m)} \\ &= \sum_{k=m^*}^{\hat{k}-1} 2^{-m^*-\log \text{BB}(m^*)+O(\log \hat{n})} \\ &= 2^{-m^*-\log \text{BB}(m^*)+O(\log \hat{n})}. \end{aligned} \quad (79)$$

The probability of this drop position, using Equation 77, is hence

$$p_S(m = m^* \mid \hat{P}, m \geq t) = \frac{2^{-m^*-\log \text{BB}(m^*)+O(\log \hat{n})}}{Z_{\text{drop}}^S} \approx \begin{cases} 1 & \text{if } m^* = t \\ 0 & \text{otherwise.} \end{cases} \quad (80)$$

Similarly, the number of strings with a specific complexity k^* , given \hat{P} and $m \geq t$, is

$$\begin{aligned} &\sum_{m=t}^{\hat{k}-1} \sum_{k=m}^{\hat{k}-1} \mathbf{1}[k = k^*] 2^{k-m+O(\log \hat{n})} 2^{-k-\log \text{BB}(m)} \\ &= \sum_{m=t}^{k^*} 2^{-m-\log \text{BB}(m)+O(\log \hat{n})} \\ &= 2^{-t-\log \text{BB}(t)+O(\log \hat{n})}. \end{aligned} \quad (81)$$

The resulting probability distribution for k^* is approximately constant across the available range:

$$p_S(k = k^* \mid \hat{P}, m \geq t) = \frac{2^{-t-\log \text{BB}(t)+O(\log \hat{n})}}{Z_{\text{drop}}^S} = 2^{O(\log \hat{n})}. \quad (82)$$

B.4.3 PROOF OF PROPOSITION 4.4(II) & (III): EXPECTED VALUES

Proof. Probabilistically weighted sums of the number of strings with a certain profile, multiplied with the respective value of m and k :

$$\begin{aligned}
M_S &:= g_{\text{drop}} \sum_{m=t}^{\hat{k}-1} \sum_{k=m}^{\hat{k}-1} m 2^{-m-\log \text{BB}(m)} \\
&\approx g_{\text{drop}} \sum_{k=t}^{\hat{k}-1} t 2^{-t-\log \text{BB}(t)} \\
&= g_{\text{drop}} t(\hat{k}-t) 2^{-t-\log \text{BB}(t)} \tag{83}
\end{aligned}$$

$$\begin{aligned}
K_S &:= g_{\text{drop}} \sum_{m=t}^{\hat{k}-1} \sum_{k=m}^{\hat{k}-1} k 2^{-m-\log \text{BB}(m)} \\
&\approx g_{\text{drop}} \sum_{k=t}^{\hat{k}-1} k 2^{-t-\log \text{BB}(t)} \\
&= g_{\text{drop}} (\hat{k}-t)(\hat{k}+t-1) 2^{-t-1-\log \text{BB}(t)} \tag{84}
\end{aligned}$$

The expected values are then

$$\mathbb{E}[m \mid \hat{P}, m \geq t] \approx \frac{M_S}{Z_{\text{drop}}^S} = t, \text{ and} \tag{85}$$

$$\mathbb{E}[k \mid \hat{P}, m \geq t] \approx \frac{K_S}{Z_{\text{drop}}^S} = \frac{\hat{k}+t-1}{2}. \tag{86}$$

Since the probability of a further drop $p_S(m \geq t \mid \hat{P}) \approx 0$, the unconditioned expectations remain at their observed values:

$$\mathbb{E}[m \mid \hat{P}] \approx \hat{m}, \text{ and } \mathbb{E}[k \mid \hat{P}] \approx \hat{k}. \tag{87}$$

Consequently, the expected compression progress is 0. \square

Following this same logic for a finite window d , both the drop probability and the expected progress vanish:

$$p_S(\text{drop within } d \mid \hat{P}) \approx 0, \text{ and } \mathbb{E}_{x \sim S}[k_{<t+\Delta c} \mid \hat{P}] \approx \hat{k}. \tag{88}$$